

## 2.3 Money Outcomes and Risk Aversion

Ref: MWG 6.C

If individual is a subjective expected utility maximizer, then  $\succsim$  over acts can be characterized by  $\pi$ , prob. measure on  $S$  representing beliefs and preference scaling utility function  $u : X \rightarrow \mathbb{R}$ .

So can identify act  $a = [\delta_{x_1}, E_1; \dots; \delta_{x_n}, E_n]$  with lottery  $L = [x_1, p_1; \dots; x_n, p_n]$  where  $p_i = \pi(E_i)$ .

Focus on situation where outcomes are amounts of wealth.

An act is now a random variable  $\tilde{x} : S \rightarrow X$ .

Identify act with its *cumulative distribution function (CDF)*

$$F_{\tilde{x}}(x) = \pi(s \in S : \tilde{x}(s) \leq x)$$

prob. realized outcome no greater than  $x$ .

Axioms place no restrictions on preference scaling utility function for wealth, but economics does.

1.  $u$  is increasing (or  $u'(x) > 0$ ).
2.  $u$  is concave (or  $u''(x) \leq 0$ )
3.  $u'''(x) > 0$  (or  $u'(\cdot)$  is convex)

1. “more is better” – local non-satiation
2. Risk aversion
3. Decreasing absolute risk aversion.

**Definition 2.3.1:** An individual is (weakly) risk averse if for any act  $\tilde{x}$ , the act that yields

$$E[\tilde{x}] = \int x dF_{\tilde{x}}(x) \left( = \sum_{x \in X} x \pi(s \in S : \tilde{x}(s) = x) \right)$$

with certainty is weakly preferred to  $\tilde{x}$ .

**Proposition 2.3.1:** If

$$U(\tilde{x}) = \int u(x) dF_{\tilde{x}}(x) \left( = \sum_{x \in X} u(x) \pi(s \in S : \tilde{x}(s) = x) \right)$$

represents  $\succsim$ , then  $\succsim$  exhibits (weak) risk aversion *if and only if* the preference-scaling utility function  $u$  is concave.

**Proof.**

**I.** concave  $u \Rightarrow$  risk aversion: By Jensen's inequality, if  $u(\cdot)$  is concave then

$$\int u(x) dF(x) \leq u\left(\int x dF(x)\right), \text{ for all } F(\cdot)$$

**II.** risk aversion  $\Rightarrow u$  is concave.

(We will show  $u$  not concave  $\Rightarrow \succsim$  does not exhibit risk aversion.)

Suppose  $u$  is not concave. That is, there exists  $y, z \in \mathbb{R}_+$  and  $\alpha \in (0, 1)$  satisfying

$$u(\alpha y + (1 - \alpha)z) < \alpha u(y) + (1 - \alpha)u(z).$$

But it then follows for the event  $E$  with  $\pi(E) = \alpha$  and the act  $\tilde{x}$ , where

$$\tilde{x}(s) = \begin{cases} y & \text{if } s \in E \\ z & \text{if } s \notin E \end{cases}, \text{ \& so } F_{\tilde{x}}(x) = \begin{cases} 0 & \text{if } x < y \\ \alpha & \text{if } x \in [y, z) \\ 1 & \text{if } x \geq z \end{cases},$$

we have

$$\begin{aligned} \int u(x) dF_{\tilde{x}}(x) &= \alpha u(y) + (1 - \alpha)u(z) > u(\alpha y + (1 - \alpha)z) \\ &= u\left(\int x dF_{\tilde{x}}(x)\right) \end{aligned}$$

## 2.4 Measures of Risk Aversion

**Certainty Equivalent** defined as

$$c(\tilde{x}, u) = u^{-1} \left( \int u(x) dF_{\tilde{x}}(x) \right)$$

**Obs:** If risk averse, then *risk premium* given by

$$\int x dF_{\tilde{x}}(x) - c(\tilde{x}, u)$$

is non-negative.

**Probability Risk Premium** Consider gamble  $\pm \varepsilon$  at base wealth  $x$ .

*Probability risk premium* implicitly defined by

$$\left( \frac{1}{2} + \pi(x, \varepsilon, u) \right) u(x + \varepsilon) + \left( \frac{1}{2} - \pi(x, \varepsilon, u) \right) u(x - \varepsilon) = u(x)$$

**Obs:** If risk averse, then  $\frac{1}{2} > \pi(x, \varepsilon, u) \geq 0$ .

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**Portfolio Problem:**  $w$  – initial wealth;  $\tilde{x}$  – random return on risky asset with CDF  $F_{\tilde{x}}(\cdot)$  and  $r$  – riskless return on safe asset.

$$\max_{\alpha \in [0,1]} \int u(\alpha wz + (1 - \alpha) wr) dF_{\tilde{x}}(z)$$

Solution  $\alpha(\tilde{x}, r, w, u)$ .

**Proposition 2.4.1** Three measures are equivalent in the sense that for two preference-scaling utility functions  $v(\cdot)$  and  $u(\cdot)$ , the following are equivalent.

1.  $c(\tilde{x}, v) \leq c(\tilde{x}, u)$  for all  $\tilde{x}$
2.  $\pi(x, \varepsilon, v) \geq \pi(x, \varepsilon, u)$  for all  $x, \varepsilon$
3.  $\alpha(\tilde{x}, r, w, v) \leq \alpha(\tilde{x}, r, w, u)$  for all  $\tilde{x}, r, w$
4.  $v$  is an *increasing* and *concave* transformation of  $u$ , that is, for some concave function  $\phi$ ,

$$v(x) = \phi(u(x))$$

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**Proof of (4)  $\Rightarrow$  (3).**

Set  $\alpha^* := \alpha(\tilde{x}, r, w, u)$  and set  $x^*(z) := \alpha^* wz + (1 - \alpha^*) wr$ .  $\alpha^*$  satisfies the FONC of portfolio problem for  $u$ . That is,

$$\int u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) = 0$$

That is,

$$-\int_{z < r} u'(x^*(z)) w(r - z) dF_{\tilde{x}}(z) + \int_{z > r} u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) = 0$$

Take the derivative of

$$\int v(w(\alpha wz + (1 - \alpha) wr)) dF(z) \equiv \int \phi(u(w(\alpha wz + (1 - \alpha) wr))) dF(z)$$

wrt  $\alpha$  and evaluate it at  $\alpha^*$ .

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This yields

$$\begin{aligned} & \int \phi'(u(x^*(z))) u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) \\ &= - \int_{z < r} \phi'(u(x^*(z))) u'(x^*(z)) w(r - z) dF_{\tilde{x}}(z) \\ & \quad + \int_{z > r} \phi'(u(x^*(z))) u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) \end{aligned}$$

As  $\phi'' \leq 0$ ,

$$\begin{aligned} \phi'(u(x^*(z))) &\geq \phi'(u(wr)) \text{ for } z < r \\ \phi'(u(x^*(z))) &\leq \phi'(u(wr)) \text{ for } z > r \end{aligned}$$

Hence

$$\begin{aligned} & \int \phi'(u(x^*(z))) u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) \\ &\leq \int \phi'(u(wr)) u'(x^*(z)) w(z - r) dF_{\tilde{x}}(z) = 0 \end{aligned}$$

So

$$\alpha(\tilde{x}, r, w, v) \leq \alpha^* = \alpha(\tilde{x}, r, w, u)$$

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## 2.4.1 Coefficient of Absolute Risk Aversion

$$r_A(x, u) \equiv -\frac{u''(x)}{u'(x)} \geq 0 \text{ if } u \text{ concave}$$

### Proposition 2.4.2

$$v(x) = \phi(u(x)), \phi' > 0, \phi'' \leq 0 \Leftrightarrow -\frac{v''(x)}{v'(x)} \geq -\frac{u''(x)}{u'(x)} \text{ for all } x.$$

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**Interpretation.** a)  $-u''(x)$  is proportional to second-order loss arising from a small fair gamble.

$$\begin{aligned} & \frac{1}{2}u(x + \varepsilon) + \frac{1}{2}u(x - \varepsilon) \\ \approx & \frac{1}{2}u(x) + \frac{1}{2}u'(x)\varepsilon + \frac{1}{4}u''(x)\varepsilon^2 + \frac{1}{2}u(x) - \frac{1}{2}u'(x)\varepsilon + \frac{1}{4}u''(x)\varepsilon^2 \end{aligned}$$

Hence

$$u(x) - \frac{1}{2}u(x + \varepsilon) + \frac{1}{2}u(x - \varepsilon) \approx -u''(x) \frac{1}{2}\varepsilon^2$$

b)  $u'(x)$  is proportional to first-order loss arising from paying a 'premium'  $d$

$$u(x - d) \approx u(x) - u'(x)d$$

I.e.

$$u(x) - u(x - d) \approx u'(x)d$$

So  $-u''(x)/u'(x)$  is proportional to the *marginal rate of substitution* between accepting a small gamble and paying a small premium.

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## 2.4.2 Hypothesis of Decreasing Absolute Risk Aversion

**DARA:**  $r_A(x, u)$  is a decreasing function of  $x$ .

**Proposition 2.4.3** The following properties are equivalent:

1. The Bernoulli utility index exhibits DARA.
2. Whenever  $y < z$ , the function  $u_y(x) \equiv u(y + x)$  is a *concave* transformation of the function  $u_z(x) \equiv u(z + x)$ .

3. For any distribution  $F$

$$\int (y + x) dF(x) - u^{-1} \left( \int (y + x) dF(x) \right)$$

is decreasing in  $y$ . That is, the higher is  $y$ , the less willing is the individual in paying a premium to get rid of the risk.

4. The probability premium  $\pi(x, \varepsilon, u)$  is decreasing in  $x$ .

5. For any  $F$ , if

$$\int u(x + \varepsilon) dF(\varepsilon) \geq u(x)$$

and  $x < y$ , then

$$\int u(y + \varepsilon) dF(\varepsilon) \geq u(y).$$

That is, the set of acceptable risks is *increasing* in  $x$ .

**Obs.**  $\frac{d}{dx}(r_A(x, u)) = \frac{d}{dx} \left( -\frac{u''(x)}{u'(x)} \right) = \frac{-u'''(x)u'(x) + (u''(x))^2}{(u'(x))^2}$

So *necessary* condition for DARA is  $u'''(x) > 0$ .

*Constant absolute risk aversion*  $u(x) = -\exp(-\alpha x)$ .

### 2.4.3 Hypothesis of Increasing Relative Risk Aversion

Recall portfolio problem and solution  $\alpha(F, r, w, u)$ .

**IRRA** says  $\alpha(F, r, w, u)$  is *decreasing* in  $w$ .

$$r_R(x, u) = -\frac{u''(x)x}{u'(x)} \text{ is increasing in } x$$

*Constant Relative Risk Aversion:*

$$u(x) = \ln x \rightarrow r_R(x, u) = 1$$

$$u(x) = \frac{1}{1-\alpha} x^{1-\alpha}, \alpha \neq 1$$

### 2.5 Mean-Variance Analysis

Justified in EU Theory if either:

1. Utility function is quadratic

$$u(x) = \begin{cases} x - bx^2/2 & \text{if } x \in [0, 1/b) \\ 1/(2b) & \text{if } x \geq 1/b \end{cases}$$

or;

2. distributions completely characterized by mean and variance.

- e.g. class of distributions are all *normally* distributed.

## Problems

1. DARA not satisfied.
2. Not all distributions normal, nor characterized just by first two moments.

Consider

$$\begin{aligned}\tilde{x} &= \begin{cases} 100 & \text{if } s \in A \\ 1 & \text{if } s \notin A \end{cases} \quad \text{with } \pi(A) = 0.2 \\ \tilde{y} &= \begin{cases} 1090 & \text{if } s \in B \\ 10 & \text{if } s \notin B \end{cases} \quad \text{with } \pi(B) = 0.01\end{aligned}$$

$$E[\tilde{x}] = 0.2 \times 100 + 0.8 \times 1 = 20.8$$

$$E[\tilde{y}] = 0.01 \times 1090 + 0.99 \times 10 = 20.8$$

$$\text{Var}(\tilde{x}) = (20.8 - 1)^2 \times 0.8 + (100 - 20.8)^2 \times 0.2 = 1,568.16$$

$$\text{Var}(\tilde{y}) = (20.8 - 10)^2 \times 0.99 + (1090 - 20.8)^2 \times 0.01 = 11,547.36$$

Say  $u(x) = \ln x$

$$U(\tilde{x}) = 0.8 \ln 1 + 0.2 \ln 100 = 0.92$$

$$U(\tilde{y}) = 0.99 \ln 10 + 0.01 \ln 1090 = 2.35$$



## 2.6 Stochastic Dominance Relations.

**DEFN:** The distribution  $F(\cdot)$  *first-order stochastically dominates*  $G(\cdot)$  if  $F(x) \leq G(x)$  for every  $x$ .

**PROPOSITION:** If  $F(\cdot)$  second-order stochastically dominates  $G(\cdot)$  then

$$\int u(x) dF(x) \geq \int u(x) dG(x)$$

for every non-decreasing  $u$ .

**DEFN:** The simple distribution  $F(\cdot)$  constitutes an *elementary first-order improvement in risk* over the simple distribution  $G(\cdot)$  if for some simple distribution  $H(\cdot)$ , some  $\alpha \in [0, 1]$ , and pair of outcomes  $x \geq y$  :

$$F(\cdot) = (1 - \alpha) H(\cdot) + \alpha \delta_x(\cdot)$$

$$G(\cdot) = (1 - \alpha) H(\cdot) + \alpha \delta_y(\cdot)$$

**FACT:** If  $F(\cdot)$  first-order stochastically dominates  $G(\cdot)$  then there exists two sequences of simple distributions,  $\langle F_n(\cdot) \rangle$  and  $\langle G_n(\cdot) \rangle$  such that

$$F_n \rightarrow F, G_n \rightarrow G$$

and for each  $n$ , there  $F_n$  can be obtained from  $G_n$  by a finite sequence of elementary first-order improvements in risk.

**DEFN:** The distribution  $F(\cdot)$  *second-order stochastically dominates*  $G(\cdot)$  if  $\int_0^x F(t) dt \leq \int_0^x G(t) dt$ .

**DEFN:** The simple distribution  $F(\cdot)$  constitutes an *elementary second-order improvement in risk* over the simple distribution  $G(\cdot)$  if for some distribution  $H(\cdot)$ , some  $\alpha, \beta \in [0, 1]$ , and for some three outcomes  $x \geq \beta y + (1 - \beta)z$ :

$$F(\cdot) = (1 - \alpha)H(\cdot) + \alpha\delta_x(\cdot)$$

$$G(\cdot) = (1 - \alpha)H(\cdot) + \alpha[\beta\delta_y(\cdot) + (1 - \beta)\delta_z(\cdot)]$$

**FACT:** If  $F(\cdot)$  second-order stochastically dominates  $G(\cdot)$  then there exists two sequences of simple distributions,  $\langle F_n(\cdot) \rangle$  and  $\langle G_n(\cdot) \rangle$  such that

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and for each  $n$ , there  $F_n$  can be obtained from  $G_n$  by a finite sequence of elementary second-order improvements in risk.

**PROPOSITION:** If  $F(\cdot)$  second-order stochastically dominates  $G(\cdot)$  then

$$\int u(x) dF(x) \geq \int u(x) dG(x)$$

for every non-decreasing and *concave*  $u$ .