Heteroskedasticity

Chapter 8

Chapter 8: Heteroskedasticity

- 8.1 The Nature of Heteroskedasticity
- 8.2 Using the Least Squares Estimator
- 8.3 The Generalized Least Squares Estimator
- 8.4 Detecting Heteroskedasticity

$$E(y) = \beta_1 + \beta_2 x$$

(8.1)

$$e_i = y_i - E(y_i) = y_i - \beta_1 - \beta_2 x_i$$

(8.2)

$$y_i = \beta_1 + \beta_2 x_i + e_i$$

(8.3)

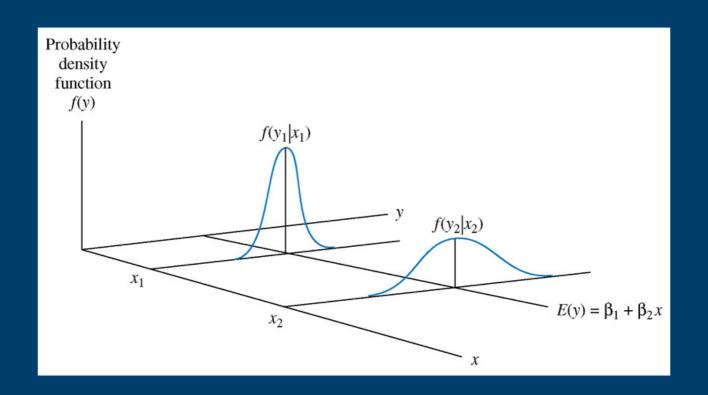


Figure 8.1 Heteroskedastic Errors

$$E(e_i) = 0$$
 $var(e_i) = \sigma^2$ $cov(e_i, e_j) = 0$

$$\operatorname{var}(y_i) = \operatorname{var}(e_i) = h(x_i)$$

$$\hat{y}_i = 83.42 + 10.21 x_i$$

$$\hat{e}_i = y_i - 83.42 - 10.21x_i$$

(8.4)

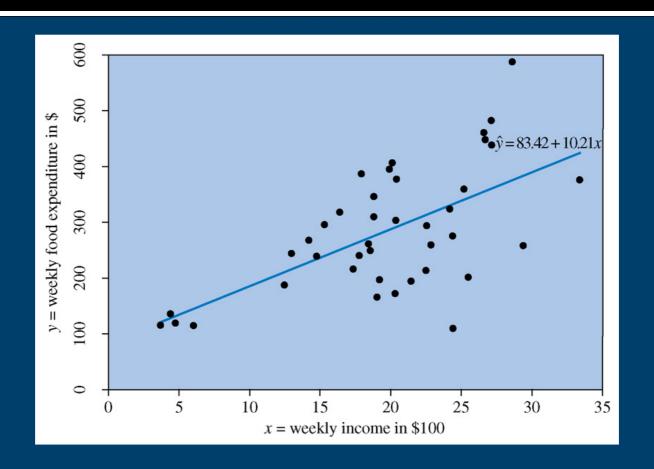


Figure 8.2 Least Squares Estimated Expenditure Function and Observed Data Points

The existence of heteroskedasticity implies:

- The least squares estimator is still a linear and unbiased estimator, but it is no longer best. There is another estimator with a smaller variance.
- The standard errors usually computed for the least squares estimator are incorrect. Confidence intervals and hypothesis tests that use these standard errors may be misleading.

$$y_i = \beta_1 + \beta_2 x_i + e_i$$
 $var(e_i) = \sigma^2$

$$\operatorname{var}(b_2) = \frac{\sigma^2}{\sum_{i=1}^{N} (x_i - \overline{x})^2}$$

(8.6)

$$y_i = \beta_1 + \beta_2 x_i + e_i$$

$$\operatorname{var}(e_i) = \sigma_i^2$$

(8.7)

$$var(b_2) = \sum_{i=1}^{N} w_i^2 \sigma_i^2 = \frac{\sum_{i=1}^{N} \left[(x_i - \overline{x})^2 \sigma_i^2 \right]}{\left[\sum_{i=1}^{N} (x_i - \overline{x})^2 \right]^2}$$
(8.8)

$$\operatorname{var}(b_2) = \sum_{i=1}^{N} w_i^2 \hat{e}_i^2 = \frac{\sum_{i=1}^{N} \left[(x_i - \overline{x})^2 \hat{e}_i^2 \right]}{\left[\sum_{i=1}^{N} (x_i - \overline{x})^2 \right]^2}$$

(8.9)

$$\hat{y}_i = 83.42 + 10.21x_i$$
(27.46) (1.81) (White se)
(43.41) (2.09) (incorrect se)

White:
$$b_2 \pm t_c \operatorname{se}(b_2) = 10.21 \pm 2.024 \times 1.81 = [6.55, 13.87]$$

Incorrect:
$$b_2 \pm t_c \operatorname{se}(b_2) = 10.21 \pm 2.024 \times 2.09 = [5.97, 14.45]$$

8.3 The Generalized Least Squares Estimator

$$y_i = \beta_1 + \beta_2 x_i + e_i$$

$$E(e_i) = 0$$
 $var(e_i) = \sigma_i^2$ $cov(e_i, e_j) = 0$

(8.10)

$$\operatorname{var}(e_i) = \sigma_i^2 = \sigma^2 x_i$$

(8.11)

$$\frac{y_i}{\sqrt{x_i}} = \beta_1 \left(\frac{1}{\sqrt{x_i}}\right) + \beta_2 \left(\frac{x_i}{\sqrt{x_i}}\right) + \frac{e_i}{\sqrt{x_i}}$$

(8.12)

$$y_i^* = \frac{y_i}{\sqrt{x_i}}$$
 $x_{i1}^* = \frac{1}{\sqrt{x_i}}$ $x_{i2}^* = \frac{x_i}{\sqrt{x_i}} = \sqrt{x_i}$ $e_i^* = \frac{e_i}{\sqrt{x_i}}$ (8.13)

$$y_i^* = \beta_1 x_{i1}^* + \beta_2 x_{i2}^* + e_i^*$$

(8.14)

$$\operatorname{var}(e_i^*) = \operatorname{var}\left(\frac{e_i}{\sqrt{x_i}}\right) = \frac{1}{x_i} \operatorname{var}(e_i) = \frac{1}{x_i} \sigma^2 x_i = \sigma^2$$
 (8.15)

To obtain the best linear unbiased estimator for a model with heteroskedasticity of the type specified in equation (8.11):

- 1. Calculate the transformed variables given in (8.13).
- 2. Use least squares to estimate the transformed model given in (8.14).

The generalized least squares estimator is as a **weighted least squares** estimator. Minimizing the sum of squared transformed errors that is given by:

$$\sum_{i=1}^{N} e_i^{*2} = \sum_{i=1}^{N} \frac{e_i^2}{x_i} = \sum_{i=1}^{N} (x_i^{-1/2} e_i)^2$$

When $\sqrt{x_i}$ is small, the data contain more information about the regression function and the observations are weighted heavily. When $\sqrt{x_i}$ is large, the data contain less information and the observations are weighted lightly.

$$\hat{y}_i = 78.68 + 10.45x_i$$
(se) (23.79) (1.39)

$$\beta_2 \pm t_c \operatorname{se}(\beta_2) = 10.451 \pm 2.024 \times 1.386 = [7.65, 13.26]$$

$$var(e_i) = \sigma_i^2 = \sigma^2 x_i^{\gamma}$$

(8.17)

$$\ln(\sigma_i^2) = \ln(\sigma^2) + \gamma \ln(x_i)$$

$$\sigma_i^2 = \exp\left(\ln(\sigma^2) + \gamma \ln(x_i)\right)$$

$$= \exp(\alpha_1 + \alpha_2 z_i)$$

(8.18)

$$\sigma_i^2 = \exp(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_s z_{iS})$$

(8.19)

$$\ln(\sigma_i^2) = \alpha_1 + \alpha_2 z_i$$

(8.20)

$$y_i = E(y_i) + e_i = \beta_1 + \beta_2 x_i + e_i$$

$$\ln(\hat{e}_i^2) = \ln(\sigma_i^2) + v_i = \alpha_1 + \alpha_2 z_i + v_i$$

$$\ln(\hat{\sigma}_i^2) = .9378 + 2.329 z_i$$

$$\hat{\sigma}_i^2 = \exp(\alpha z_i + \alpha_1 z_i)$$

$$\left(\frac{y_i}{\sigma_i}\right) = \beta_1 \left(\frac{1}{\sigma_i}\right) + \beta_2 \left(\frac{x_i}{\sigma_i}\right) + \left(\frac{e_i}{\sigma_i}\right)$$

(8.21)

$$\operatorname{var}\left(\frac{e_i}{\sigma_i}\right) = \left(\frac{1}{\sigma_i^2}\right) \operatorname{var}(e_i) = \left(\frac{1}{\sigma_i^2}\right) \sigma_i^2 = 1$$

$$y_i^* = \left(\frac{y_i}{y_i}\right)$$

$$y_i^* = \left(\frac{y_i}{\sigma_i}\right) \qquad x_{i1}^* = \left(\frac{1}{\sigma_i}\right) \qquad x_{i2}^* = \left(\frac{x_i}{\sigma_i}\right)$$

$$y_i^* = \beta_1 x_{i1}^* + \beta_2 x_{i2}^* + e_i^*$$

(8.24)

$$y_i = \beta_1 + \beta_2 x_{i2} + \dots + \beta_k x_{iK} + e_i$$

$$var(e_i) = \sigma_i^2 = exp(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_s z_{iS})$$

(8.26)

The steps for obtaining a feasible generalized least squares estimator for $\beta_1, \beta_2, ..., \beta_K$ are:

- 1. Estimate (8.25) by least squares and compute the squares of the least squares residuals \hat{e}_i^2 .
- 2. Estimate $\alpha_1, \alpha_2, ..., \alpha_S$ by applying least squares to the equation $\ln \hat{e}_i^2 = \alpha_1 + \alpha_2 z_{i2} + \cdots + \alpha_S z_{iS} + v_i$

- 3. Compute variance estimates $\hat{\sigma}_i^2 = \exp(\alpha z_1 + \alpha_2 z_{i2} + \cdots + \alpha_s z_{is})$
- 4. Compute the transformed observations defined by (8.23), including $x_{i3}^*, \dots, x_{iK}^*$ if K > 2.
- 5. Apply least squares to (8.24), or to an extended version of (8.24) if K > 2.

$$\hat{y}_i = 76.05 + 10.63x$$
(se) (9.71) (.97)

$$WAGE = -9.914 + 1.234EDUC + .133EXPER + 1.524METRO$$
 (se) (1.08) (.070) (.015) (.431)

$$WAGE_{Mi} = \beta_{M1} + \beta_2 EDUC_{Mi} + \beta_3 EXPER_{Mi} + e_{Mi}$$
 $i = 1, 2, ..., N_M$ (8.29a)

$$WAGE_{Ri} = \beta_{R1} + \beta_2 EDUC_{Ri} + \beta_3 EXPER_{Ri} + e_{Ri}$$
 $i = 1, 2, ..., N_R$ (8.29b)

$$b_{M1} = -9.914 + 1.524 = -8.39$$

(8.28)

$$var(e_{Mi}) = \sigma_M^2 \qquad var(e_{Ri}) = \sigma_R^2$$
 (8.30)

$$\sigma_{M}^{2} = 31.824$$
 $\sigma_{R}^{2} = 15.243$

$$b_{M1} = -9.052$$
 $b_{M2} = 1.282$ $b_{M3} = .1346$

$$b_{R1} = -6.166$$
 $b_{R2} = .956$ $b_{R3} = .1260$

$$\left(\frac{WAGE_{Mi}}{\sigma_{M}}\right) = \beta_{M1} \left(\frac{1}{\sigma_{M}}\right) + \beta_{2} \left(\frac{EDUC_{Mi}}{\sigma_{M}}\right) + \beta_{3} \left(\frac{EXPER_{Mi}}{\sigma_{M}}\right) + \left(\frac{e_{Mi}}{\sigma_{M}}\right)$$

$$i = 1, 2, \dots, N_{M}$$

$$\left(\frac{WAGE_{Ri}}{\sigma_R}\right) = \beta_{R1} \left(\frac{1}{\sigma_R}\right) + \beta_2 \left(\frac{EDUC_{Ri}}{\sigma_R}\right) + \beta_3 \left(\frac{EXPER_{Ri}}{\sigma_R}\right) + \left(\frac{e_{Ri}}{\sigma_R}\right)$$

$$i = 1, 2, \dots, N_R$$

(8.31b)

Feasible generalized least squares:

1. Obtain estimated $\hat{\sigma}_M$ and $\hat{\sigma}_R$ by applying least squares separately to the metropolitan and rural observations.

2.
$$\hat{\sigma}_i = \begin{cases} \hat{\sigma}_M & \text{when } METRO_i = 1 \\ \hat{\sigma}_R & \text{when } METRO_i = 0 \end{cases}$$

3. Apply least squares to the transformed model

$$\left(\frac{WAGE_{i}}{C_{i}}\right) = \beta_{R1}\left(\frac{1}{\sigma_{i}}\right) + \beta_{2}\left(\frac{EDUC_{i}}{C_{i}}\right) + \beta_{3}\left(\frac{EXPER_{i}}{\sigma_{i}}\right) + \delta\left(\frac{METRO_{i}}{C_{i}}\right) + \left(\frac{e_{i}}{\sigma_{i}}\right)$$
(8.32)

$$WAGE = -9.398 + 1.196EDUC + .132EXPER + 1.539METRO$$
(se) (1.02) (.069) (.015) (.346)

Remark: To implement the generalized least squares estimators described in this Section for three alternative heteroskedastic specifications, an assumption about the form of the heteroskedasticity is required. Using least squares with White standard errors avoids the need to make an assumption about the form of heteroskedasticity, but does not realize the potential efficiency gains from generalized least squares.

8.4.1 Residual Plots

- Estimate the model using least squares and plot the least squares residuals.
- With more than one explanatory variable, plot the least squares residuals against each explanatory variable, or against \hat{y}_i , to see if those residuals vary in a systematic way relative to the specified variable.

8.4.2 The Goldfeld-Quandt Test

$$F = \frac{\hat{\sigma}_M^2 / \sigma_M^2}{\hat{\sigma}_R^2 / \sigma_R^2} \square F_{(N_M - K_M, N_R - K_R)}$$

 $H_0: \sigma_M^2 = \sigma_R^2$ against $H_0: \sigma_M^2 \neq \sigma_R^2$

(8.35)

(8.34)

$$F = \frac{\hat{\sigma}_M^2}{\hat{\sigma}_R^2} = \frac{31.824}{15.243} = 2.09$$

8.4.2 The Goldfeld-Quandt Test

$$\hat{\sigma}_1^2 = 3574.8$$

$$\hat{\sigma}_{2}^{2} = 12,921.9$$

$$F = \frac{\hat{\sigma}_2^2}{\hat{\sigma}_1^2} = \frac{12,921.9}{3574.8} = 3.61$$

8.4.3 Testing the Variance Function

$$y_i = E(y_i) + e_i = \beta_1 + \beta_2 x_{i2} + \dots + \beta_K x_{iK} + e_i$$

(8.36)

$$var(y_i) = \sigma_i^2 = E(e_i^2) = h(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS})$$

(8.37)

$$h(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS}) = \exp(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS})$$

$$h(\alpha_1 + \alpha_2 z_i) = \exp\left(\ln(\sigma^2) + \gamma \ln(x_i)\right)$$

8.4.3 Testing the Variance Function

$$h(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS}) = \alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS}$$
 (8.38)

$$h(\alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS}) = h(\alpha_1)$$

$$H_0: \alpha_2 = \alpha_3 = \cdots = \alpha_S = 0$$

 H_1 : not all the α_s in H_0 are zero

(8.39)

8.4.3 Testing the Variance Function

$$var(y_i) = \sigma_i^2 = E(e_i^2) = \alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS}$$

(8.40)

$$e_i^2 = E(e_i^2) + v_i = \alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS} + v_i$$

(8.41)

$$\hat{e}_i^2 = \alpha_1 + \alpha_2 z_{i2} + \dots + \alpha_S z_{iS} + v_i$$

(8.42)

$$\chi^2 = N \times R^2 \square \chi^2_{(S-1)}$$

(8.43)

8.4.3a The White Test

$$E(y_i) = \beta_1 + \beta_2 x_{i2} + \beta_3 x_{i3}$$

$$z_2 = x_2$$
 $z_3 = x_3$ $z_4 = x_2^2$ $z_5 = x_3^2$

8.4.3b Testing the Food Expenditure Example

$$SST = 4,610,749,441$$
 $SSE = 3,759,556,169$

$$R^2 = 1 - \frac{SSE}{SST} = .1846$$

$$\chi^2 = N \times R^2 = 40 \times .1846 = 7.38$$

$$\chi^2 = N \times R^2 = 40 \times .18888 = 7.555$$
 p-value = .023

Keywords

- Breusch-Pagan test
- generalized least squares
- Goldfeld-Quandt test
- heteroskedastic partition
- heteroskedasticity
- heteroskedasticity-consistent standard errors
- homoskedasticity
- Lagrange multiplier test
- mean function
- residual plot
- transformed model
- variance function
- weighted least squares
- White test

Chapter 8 Appendices

- Appendix 8A Properties of the Least Squares
 Estimator
- Appendix 8B Variance Function Tests for Heteroskedasticity

$$y_i = \beta_1 + \beta_2 x_i + e_i$$

$$E(e_i) = 0$$
 $var(e_i) = \sigma_i^2$ $cov(e_i, e_j) = 0$ $(i \neq j)$

$$b_2 = \beta_2 + \sum w_i e_i \tag{8A.1}$$

$$w_i = \frac{x_i - \overline{x}}{\sum (x_i - \overline{x})^2}$$

$$E(b_2) = E(\beta_2) + E(\sum w_i e_i)$$
$$= \beta_2 + \sum w_i E(e_i) = \beta_2$$

$$\operatorname{var}(b_{2}) = \operatorname{var}\left(\sum w_{i}e_{i}\right)$$

$$= \sum w_{i}^{2} \operatorname{var}(e_{i}) + \sum_{i \neq j} \sum w_{i}w_{j} \operatorname{cov}\left(e_{i}, e_{j}\right)$$

$$= \sum w_{i}^{2} \sigma_{i}^{2}$$

$$= \frac{\sum \left[\left(x_{i} - \overline{x}\right)^{2} \sigma_{i}^{2}\right]}{\left[\sum \left(x_{i} - \overline{x}\right)^{2}\right]^{2}}$$
(8A.2)

$$var(b_2) \neq \frac{\sigma^2}{\sum (x_i - \overline{x})^2}$$
 (8A.3)

$$\hat{e}_{i}^{2} = \alpha_{1} + \alpha_{2}z_{i2} + \dots + \alpha_{S}z_{iS} + v_{i}$$
 (8B.1)

$$F = \frac{(SST - SSE)/(S-1)}{SSE/(N-S)}$$
(8B.2)

$$SST = \sum_{i=1}^{N} \left(\underbrace{e^2}_{i} - \overline{e^2} \right)^2 \quad \text{and} \quad SSE = \sum_{i=1}^{N} \widehat{v}_i^2$$

$$\chi^2 = (S-1) \times F = \frac{SST - SSE}{SSE / (N-S)} \square \chi^2_{(S-1)}$$
 (8B.3)

$$\operatorname{var}(e_i^2) = \operatorname{var}(v_i) = \frac{SSE}{N - S}$$
 (8B.4)

$$\chi^2 = \frac{SST - SSE}{\operatorname{Var}(e_i^2)}$$
 (8B.5)

$$\chi^2 = \frac{SST - SSE}{2\hat{\sigma}_e^4} \tag{8B.6}$$

$$\operatorname{var}\left(\frac{e_i^2}{\sigma_e^2}\right) = 2 \qquad \frac{1}{\sigma_e^4} \operatorname{var}(e_i^2) = 2 \qquad \operatorname{var}(e_i^2) = 2\sigma_e^4$$

$$var(e_i^2) = \frac{1}{N} \sum_{i=1}^{N} (e_i^2 - \overline{e^2})^2 = \frac{SST}{N}$$
 (8B.7)

$$\chi^{2} = \frac{SST - SSE}{SST / N}$$

$$= N \times \left(1 - \frac{SSE}{SST}\right)$$

$$= N \times R^{2}$$
(8B.8)